

Approximation of mean and variance for functions of random variables.

Example. Measure voltage, U , and current, I , and want to calculate the resistance $R = \frac{U}{I}$. What about $E[R]$ and $\text{Var}[R]$.

$$\text{Let } E[U] = \mu_1 \text{ and } E[I] = \mu_2$$

From a Taylor-expansion of $f(x, y) = \frac{x}{y}$ around (μ_1, μ_2) we get.

$$\begin{aligned} f(x, y) &= f(\mu_1, \mu_2) + \frac{\partial f}{\partial x} \Big|_{(\mu_1, \mu_2)} (x - \mu_1) + \frac{\partial f}{\partial y} \Big|_{(\mu_1, \mu_2)} (y - \mu_2) + \varepsilon \\ &= \frac{\mu_1}{\mu_2} + \frac{1}{\mu_2} (x - \mu_1) - \frac{\mu_1}{\mu_2^2} (y - \mu_2) + \varepsilon \end{aligned}$$

i.e. if ε is small, R can be approximated by

$$R = \frac{U}{I} \approx \frac{\mu_1}{\mu_2} + \frac{1}{\mu_2} (U - \mu_1) - \frac{\mu_1}{\mu_2^2} (I - \mu_2)$$

such that.

$$E[R] \approx \frac{\mu_1}{\mu_2} + \frac{1}{\mu_2} E[U - \mu_1] - \frac{\mu_1}{\mu_2^2} E[I - \mu_2] = \frac{\mu_1}{\mu_2},$$

and

$$\text{Var}[R] \approx \left[\frac{1}{\mu_2} \right]^2 \text{Var}[U] + \frac{\mu_1^2}{\mu_2^4} \text{Var}[I] - \frac{2\mu_1}{\mu_2^3} \text{Cov}[U, I]$$

If the uncertainty in U and I ~~is an~~ are measurement errors it is reasonable that $\text{Cov}[U, I] = 0$

In general if $Y = f(x_1, \dots, x_m)$, we have:

$$f(x_1, \dots, x_m) = f(\mu_1, \dots, \mu_m) + \sum_{j=1}^m \frac{\partial f}{\partial x_j} \Big|_{(\mu_1, \dots, \mu_m)} (x_j - \mu_j) + \varepsilon$$

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and if higher order terms are small, we have

$$E[Y] \approx f(\mu_1, \dots, \mu_n)$$

$$\text{Var}[Y] \approx \sum_{i=1}^n \left(\frac{\partial f(\mu_1, \dots, \mu_n)}{\partial x_i} \right)^2 \text{Var}[x_i] + 2 \sum_{j < k} \frac{\partial f(\mu_1, \dots, \mu_n)}{\partial x_j} \frac{\partial f(\mu_1, \dots, \mu_n)}{\partial x_k} \text{Cov}(x_j, x_k)$$

Example $U \sim N(100, 2^2)$ $R = \frac{U}{I}$
 $I \sim N(20, 1)$

We get, $E[R] \approx 5$ and $\text{Var}[R] \approx \left(\frac{1}{20} \right)^2 \cdot 4 + \left(\frac{100}{400} \right)^2 \cdot 1$

$$= \frac{4}{400} + \frac{1}{16} = \frac{1}{100} + \frac{1}{16} = 0.0775$$

$$\left. \begin{array}{l} U \sim N(100, 5^2) \\ I \sim N(20, 2^2) \end{array} \right\} \Rightarrow \begin{array}{l} E[R] \approx 5 \\ \text{Var}[R] \approx \left(\frac{1}{20} \right)^2 \cdot 25 + \left(\frac{100}{400} \right)^2 \cdot 4 \approx \underline{0.31} \end{array}$$

Purpose: Find relationships between one response variable and one or more explanatory variables. Regression analysis is a modelling tool.

11.2. Simple linear regression (linear in the coefficients)

$$\text{Model: } Y_i = \alpha + \beta X_i + \epsilon_i \quad \left. \begin{array}{l} \text{independent} \\ E[\epsilon_i] = 0, \text{Var}[\epsilon_i] = \sigma^2 \end{array} \right\}$$

$\nearrow$ response
 dependent variable

$\nwarrow$ regression variable
 independent variable
 explanatory variable

X_i may be a transformation of other variables.

i.e. $X_i = \ln(t_i)$, $X_i = \sin(t_i)$ and so on.

Observe: $(y_1, x_1), (y_2, x_2), \dots, (y_m, x_m)$ which ~~are~~ have to fulfill,

$$\begin{aligned} y_1 &= \alpha + \beta x_1 + \epsilon_1 \\ y_2 &= \alpha + \beta x_2 + \epsilon_2 \\ &\vdots \\ y_m &= \alpha + \beta x_m + \epsilon_m \end{aligned} \quad \text{or} \quad \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{bmatrix} = \begin{bmatrix} 1 & x_1 \\ 1 & x_2 \\ \vdots & \vdots \\ 1 & x_m \end{bmatrix} \begin{bmatrix} \alpha \\ \beta \end{bmatrix} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_m \end{bmatrix}$$

We need to find estimate, (a, b) , for (α, β) .

Method of Least Squares

Find the a and b that minimize

$$Q = \sum_{i=1}^m \epsilon_i^2 = \sum_{i=1}^m (y_i - a - b x_i)^2$$

$$\frac{\partial Q}{\partial a} = 0 \iff -2 \sum_{i=1}^n (y_i - a - bx_i) = 0$$

$$\frac{\partial Q}{\partial b} = 0 \iff -2 \sum_{i=1}^n (y_i - a - bx_i) x_i = 0$$

This gives the 2 normal equations

$$\sum_{i=1}^n a + \sum_{i=1}^n bx_i = \sum_{i=1}^n y_i \quad (1)$$

$$\sum_{i=1}^n ax_i + \sum_{i=1}^n bx_i^2 = \sum_{i=1}^n x_i y_i \quad (2)$$

From (1) we get:

$$na + b \sum_{i=1}^n x_i = \sum_{i=1}^n y_i \text{ which implies } a = \bar{y} - b\bar{x}$$

and substitution of a with $\bar{y} - b\bar{x}$ in (2) gives

$$(\bar{y} - b\bar{x}) \sum_{i=1}^n x_i + b \sum_{i=1}^n x_i^2 = \sum_{i=1}^n x_i y_i$$

$$\Rightarrow b \left(\sum_{i=1}^n x_i^2 - n\bar{x}^2 \right) = \sum_{i=1}^n x_i (y_i - \bar{y})$$

$$\Rightarrow b = \frac{\sum_{i=1}^n x_i (y_i - \bar{y})}{\sum_{i=1}^n x_i^2 - n\bar{x}^2} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\text{Here is used: } \sum_{i=1}^n \bar{x} (y_i - \bar{y}) = \bar{x} \sum_{i=1}^n (y_i - \bar{y}) = 0$$

$$\sum_{i=1}^n (x_i - \bar{x})^2 = \sum_{i=1}^n (x_i^2 - 2x_i\bar{x} + \bar{x}^2) = \sum_{i=1}^n x_i^2 - 2n\bar{x}^2 + n\bar{x}^2$$

$$= \sum_{i=1}^n x_i^2 - n\bar{x}^2$$

$$\sum_{i=1}^n \bar{y} (x_i - \bar{x}) = \bar{y} \sum_{i=1}^n (x_i - \bar{x}) = 0$$

The corresponding estimators are: $\hat{\beta} = \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}$

and $\hat{\alpha} = \bar{y} - \hat{\beta} \bar{x}$.

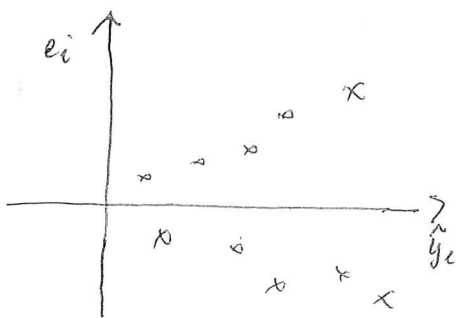
The estimated expected regression line (estimated model) is:

$$\hat{y} = a + bx$$

Residuals: $e_i = y_i - \hat{y}_i = y_i - a - bx_i$

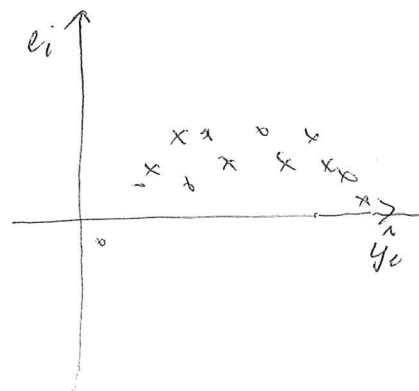
These should be plotted against $\hat{y}_i$, against x_i and eventually other regression variables. Normal-plot should be used to check for normal distribution.

Pattern in the residuals indicates that the model does not fit.



Variance not constant.

Transform y

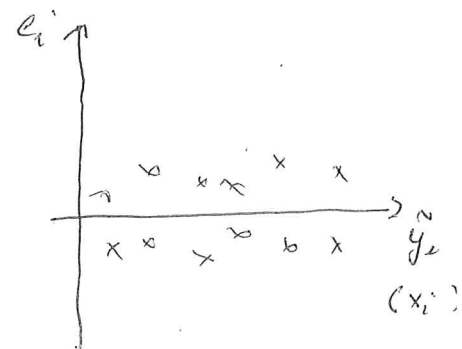


Should consider

second order

terms

$$E[y_i] = \alpha + \beta_1 x_i + \beta_2 x_i^2$$



No pattern

(ok.)